SELECTION EFFECTS OF REPEATABILITY CRITERIA APPLIED TO LUNG SPIROMETRY

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Elsen E. A. (Harvard School of Public Health, Boston, MA 02115), J. M. Robins, I. A. Greaves and D. H. Wegman. Selection effects of repeatability criteria applied to lung splrometry. Am J Epidemiol 1984; 120:734-42.

The potential for introducing bias In studies of pulmonary function by the exclusion of subjects with nonrepeatable measurements was examined In a cohort of Vermont granite workers followed for five years. At each annual survey, a "test failure" was defined as a test In which the two largest forced expiratory volumes In one second (FEV,) differed by more than 200 ml. "Persistent test failure" was defined In terms of 1) the number of test failures for each worker over the six surveys and 2) the difference between the two best efforts at each survey, averaged over all surveys for each worker. The rate of FEV, loss was estimated for each subject based only on repeatable measurements. It Is widespread practice to exclude subjects from analysis who do not perform repeatable lung function tests. The authors found that subjects with persistent test failure were losing FEV, at a faster rate than subjects without. The results suggest that the application of rigid repeatability criteria may bias epidemlologic findings by the exclusion of many subjects with accelerated loss of lung function.

epidemlologic methods; longitudinal studies; lung volume measurements

the American Thoracic Society proposed tions and performed the test with a a series of standards for spirometry (1). smooth continuous exhalation, with ap-Later that same year, similar standards parent maximal effort, with a good start, were published by the Epidemiology and without coughing, early termination, Standardization Project (2). The Snow- glottis closure, a leak, or an obstructed bird Report included instrument specifi- mouthpiece. The participants at the cations as well as methods for the stan- Snowbird workshop recommended a mindardized measurement of forced expira- imum of three acceptable forced expiratory volumes. An acceptable curve was to tory maneuvers. In an attempt to ensure be defined by the technician's observation subject cooperation and maximal efforts,

In 1978, the Snowbird Conference of that the subject understood the instructhey further proposed that the best two of Received for publication October 17, 1983, and in three acceptable curves should not vary
final form March 1, 1984. final form March 1, 1984. by more than 100 ml or 5 per cent, which ever was greater

Most epidemiologic studies of pulmo-² Center for the Health Sciences, UCLA School of nary function reported after 1978 have Center for the Health Sciences, UCLA School of **nary function reported after 1978 have**
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Reprise the Partial Angeles, Company conformed with the recommended stan-Reprint requests to Dr. Ellen A. Eisen, Occupa- conformed with the recommended stan-
onal Health Program, Harvard School of Public dards, although there are differences in ealth, 665 Huntington Avenue, Boston, MA 02115. the degree of variability tolerated be-
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"100-ml or 5 per cent" rule found in the literature include a simple 100-ml or 200 ml rule, a 3 per cent rule, and a 5 per cent criterion (3-10). Pulmonary function tests that fail to satisfy whatever criterion is applied are generally regarded as unreliable and are excluded from the subsequent analysis of the measurements.

The need for a standardized method to exclude submaximal values from the analysis of pulmonary function data motivated the American Thoracic Society recommendations. To the extent that satisfactory repeatability can only be achieved for true maximal efforts, the exclusion of nonrepeatable tests would certainly reduce the risk of falsely low test results. If lung disease were sometimes responsible for certain subjects being unable to perform repeatable tests, however, the widespread adoption of spirometry standards that include repeatability criteria could result in the exclusion of data from a proportion of subjects with pulmonary disease. If the application of repeatability criteria leads to the selective exclusion of diseased individuals, their application may thus be an important source of bias in epidemiologic studies that consider the possible effects of environmental agents on the lungs.

Clinical experience suggests that subjects with airflow obstruction and pulmonary disability do not have highly repeatable lung function tests. This is particularly true for asthmatics in whom repeated forced expiratory maneuvers may exacerbate airflow narrowing (11), and for subjects with other forms of airway obstruction in whom serial measurements over several minutes may show a sequential decrease in forced expiratory volumes. It is important to note that these clinical observations relate to patients who have overt disease and are usually disabled; it is presently unknown whether "early" or mild disease might similarly influence forced expiratory volumes.

To evaluate the influence of current spirometric standards that include a repeatability criterion, we have examined the rate of decline in lung function among subjects who satisfy the repeatability criterion and among those who do not. The subjects come from a cohort of Vermont granite workers who were recruited in 1970; all were currently employed at the time of recruitment and none could be considered to have disabling lung disease at the beginning of the study.

We shall use the term *repeatability* to denote the ability of an individual at a single sitting to perform a series of expiratory maneuvers within a specified narrow range. *Test failure* will be used to describe a series of forced expirations performed on a single occasion which do not satisfy the repeatability criterion.

METHODS

Study design

An industry-wide silicosis prevention program was initiated in the Vermont granite sheds in the late 1930s. At that time, dust controls were installed in all the granite sheds which substantially reduced silica exposures, and a medical monitoring system was started. There was an initial emphasis on radiographic examinations, but in 1969, pulmonary function tests and respiratory questionnaires were added to the annual surveys. This longitudinal study consists of data collected from 1970 to 1975 in six annual medical surveys of the workers.

The population eligible for this study were the white male workers actively employed in 1970 who started work in the industry after the reduction of dust levels (after 1940). Eligibility also required that the workers were at least 25 years of age in 1970 and that they attended the medical survey in that year. After excluding those with histories of employment in other dusty trades, 713 workers were available for study.

Pulmonary function tests

Forced expiratory volume in one second $(FEV₁)$ was measured annually for all the participants in the medical surveys. A single Stead-Wells spirometer was used for all lung function testing, and the same technician administered all tests. For each subject, a minimum of three and a maximum of five measurements were recorded. A *test failure* for a given worker was defined when the two largest $FEV₁S$ differed by more than 200 ml during a particular test session (7). For all repeatable $FEV₁₈$, the mean of the two best values was used. An additional measure of repeatability was defined as the difference between the two largest acceptable FEV₁s in each session (FEV₁</sub> difference).

Two measures of persistent test failure were also defined: 1) the total number of test failures for each worker over the six medical surveys and 2) the average $FEV₁$ difference over all surveys for each worker. The latter measurement will be referred to as the *mean difference.*

Calculation of rate of FEV\ decline

An estimate of rate of $FEV₁$ decline was computed for all workers who had repeatable $FEV₁$ measurements on at least two occasions. The rate of $FEV₁$ decline for each worker was estimated in a separate regression equation in which $FEV₁$ was regressed on follow-up time as follows:

$$
FEV(i,t) = a(i) - b(i)t, \qquad (1)
$$

where $a(i)$ is the estimated FEV_1 level of the ith worker in 1970, *b(i)* is the estimate of the rate of $FEV₁$ decline for the *i*th worker, and *t* is the duration of follow-up for each measurement (i.e., $t = 0-5$ for measurements from 1970 to 1975, respectively). $\text{FEV}(i,t)$ is the mean of the two largest FEV₁s for worker *i* at time *t*. The value for $FEV(i,t)$ was used only if the repeatability criterion was satisfied; otherwise, *FEV(i,t)* was treated as a missing observation.

Exposure

Lifetime silica exposure was estimated for each subject by combining his work history in the industry with extensive industrial hygiene data collected during the study period. Lifetime exposure for each individual was defined as a sum of the dust levels in each job worked, weighted by the number of years spent in that job. The derivation of exposure estimates is described in detail in a separate report (12). On average, the dust levels were approximately half the current threshold limit value for occupational exposures to silica-containing dust.

Analytic method

Neither the rate of $FEV₁$ decline nor the $FEV₁$ difference was normally distributed. Both were skewed and had a higher proportion of observations in the tails than Gaussian distributions. Furthermore, the standard error of the mean rate of $FEV₁$ decline varied by threefold between relevant subgroups of the cohort. In order to use the rate of FEV_1 decline as the outcome in a multivariate parametric model, a logarithmic transformation was needed to stabilize the variance and to symmetrize the distribution. A constant was added to each measurement before the transformation to avoid undefined values (13).

A matrix of Spearman correlations, rather than parametric correlations, was computed between $FEV₁$ differences in each survey. A simultaneous test of whether the off-diagonal elements were nonzero was performed using an approximation of the Bartlett statistic (14).

RESULTS

Study population

The rate of $FEV₁$ decline could not be calculated for 56 workers who had attended only one survey nor for 39 who failed to perform repeatable $FEV₁s$ on at least two occasions. The mean characteristics of the remaining 618 subjects are shown in table 1. The workers had an average baseline FEV_1 which was 95 per cent of that predicted for their age and height, using a standard equation based on healthy, asymptomatic white males (15). Their mean rate of $FEV₁$ decline was 51 ml/vear.

The cohort was divided into "dropouts" (those who left during the study) and "survivors." Those who left the industry were slightly older and had a lower baseline $FEV₁$ (per cent predicted). The mean difference was approximately equal in the two groups, although its variance was greater among the dropouts. The dropouts also had a greater rate of $FEV₁$ decline than the survivors: 69 ml/year versus 48 ml/year, respectively. The proportion of current smokers was similar in the two groups.

Because of the restriction imposed on the date of hire (1940), the cohort was relatively young, and only 11 of the 103 dropouts left because of retirement. The 11 retirees were found to be losing 42 ml/ year; this rate of decline was less than that of workers who quit before retirement age. Additional data on the dropouts have been published in a separate report (16). The remainder of this analysis concentrates on those 515 workers who remained actively employed throughout the entire study period.

The overall test failure rate throughout the study was 11 per cent using the 200 ml repeatability criterion. This rate did not vary significantly between surveys (table 2). We first examined whether test failure was equally likely for all subjects at each testing session. A goodness of fit test showed that the variability in test failure rates was greater than expected

under the binomial assumption $(p =$ 0.001). Thus, a greater proportion of workers was found to have multiple test failures than would have been expected if test failures had occurred randomly, indicating that some individuals had persistent test failure (table 3).

The associations between degrees of repeatability among surveys were also illustrated by the Spearman correlation matrix of $FEV₁$ differences in each survey for each worker. Although small, the strength of the correlations increased (from $r = 0.03$ to $r = 0.14$) as the number of years between surveys decreased (from five to one). All the correlations between $FEV₁$ differences from survey to survey were positive and simultaneously greater than zero $(p < 0.001)$. This suggests that workers who performed a poorly repeatable $FEV₁$ at one survey were more likely to perform poorly repeatable tests at subsequent surveys.

Having demonstrated that persistent

	Dropouts	Survivors	Total	
No. of workers	103	515	618	
Age (years) $*$	±1.0 42.7	± 0.4 39.7	± 0.4 40.2	
$FEV1$ ⁺ level (% predicted) [*]	$±$ 1.6% 92%	± 0.7% 96%	± 0.6% 95%	
FEV , mean difference (ml)*	\pm 6.6 93.5	90.0 ± 2.5	± 2.3 90.4	
FEV ₁ decrement (ml/year)*	± 12.6 69.4	$+3.9$ 47.8	± 3.9 51.4	
Silica exposure $(mg/m^3 \times \text{years})^*$	0.773 ± 0.060	0.703 ± 0.024	0.714 ± 0.023	
% never smoked	18	17	17	
% current smokers	58	60	60	

TABLE 1 *Characteristics of the cohort of Vermont granite workers: dropouts and survivors*

* Values are means ± standard errors,

t Forced expiratory volume in one second.

		Survey no.				
No. of participants	542	477	452	438	429	446
$FEV1$ * difference (ml)†	92 ± 5	94 ± 5	94 ± 5	94 ± 5	96 ± 6	94 ± 6
$FEV1$ (% predicted)	96	95	94	94	94	94
Test failure rate	0.10	0.12	0.11	0.11	0.11	0.11

TABLE 2 *Survey effects; study of Vermont granite workers*

* Forced expiratory volume in one second.

 \dagger Values are means \pm standard errors.

test failure does occur in some individuals, we proceeded to examine the more critical question of whether the subjects with persistent test failure were a random subset of the cohort with respect to the rate of $FEV₁$ decline.

Characteristics of subjects with persistent test failure

Because participation in the medical surveys was voluntary, some workers did not attend all the surveys. There was no correlation, however, between the mean rates of $FEV₁$ decline and the number of surveys attended $(r = 0.04, p = 0.31)$. Therefore, in spite of the unequal number of opportunities for failure, the relationship between persistent test failure and the rate of FEV_1 decline was unconfounded by the degree of participation. As shown in table 3, 62 per cent of the cohort had no test failures, 25 per cent had only

one, and 12 per cent had more than one test failure; the corresponding mean rates of $FEV₁$ decline for the three groups were 46, 37, and 81 ml/year, respectively.

The subjects with persistent test failure, as defined by at least two test failures, had greater variability in their rates of FEV_1 decline than subjects without persistent failure. This was probably because of the heterogeneity of the population of individuals with persistent failure and the fewer number of test successes per individual available for the estimation of their rates of decline. When the transformed rate was treated as the outcome in a regression analysis, the association between persistent test failure and rate of $FEV₁$ decline was significant $(p = 0.05)$ after controlling for age, height, silica exposure, and current smoking status.

As a more sensitive indicator of persis-

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Frequency of test failure among survivors in the six annual surveys of Vermont granite workers

* Values are means ± standard errors.

t Forced expiratory volume in one second.

tent test failure, the differences between the two largest $FEV₁s$ at each session were averaged over all the surveys for each worker. As the mean difference increased, so did the rate of $FEV₁$ decline. This relationship is illustrated in figure 1.

 $FEV₁$ differences were found to be correlated between successive surveys; that is, an individual's $FEV₁$ differences at surveys adjacent to test failures were likely to be larger than those for the surveys that were preceded and followed by test successes. Thus, inclusion of $FEV₁$ measurements from surveys adjacent to test failures could influence estimates of the rates of $FEV₁$ decline. If test successes occurred in the early surveys and were followed by test failures, the last valid $FEV₁$ would more likely underestimate the "true" FEV_1 than the earlier measurements, thereby causing the rate of $FEV₁$ decline to be overestimated. Conversely, test failures in the early surveys followed by successes could result in underestimation of the rate of $FEV₁$ decline.

To assess the impact of these possible biases on our results, the juxtaposition of test failures and successes was examined among subjects with a mean difference of at least 200 ml. Without knowledge of the calculated rates of $FEV₁$ decline, two of the authors (E.A.E. and J.M.R.) independently divided subjects into three groups according to the relative ordering of their failures and successes: 1) those with test patterns likely to cause overestimation of *FEV1* decline (successes followed by failures); 2) those likely to be unbiased (successes and failures interspersed); and 3) those likely to cause underestimation of FEV_1 decline (failures followed by successes). The means of the observed rates of $FEV₁$ decline for these three groups were in the anticipated order: 171 ml/year, 99 ml/year, and 67 ml/ year, respectively. It is important to note that even subjects in groups 2 and 3 (with unbiased or underestimated rates of $FEV₁$ decline) had, on average, greater rates of FEV_1 decline than the members of the cohort who had a mean difference

FIGURE 1. The average rates of FEV, loss are shown for individuals grouped by the following intervals of average FEV, difference: <50 ml, 50-100 ml, 100-150 ml, 150-200 ml, 200-250 ml, 250-300 ml, >300 ml. The number of subjects in each interval is indicated.

of less than 200 ml (44 ml/year). Therefore, the increased rate of $FEV₁$ decline observed among subjects with persistent test failure cannot be explained by a systematic pattern of test failures and successes causing an overestimation of their rates of $FEV₁$ declines.

DISCUSSION

Whenever an individual performs an $FEV₁$ maneuver, there exists a hypothetic, "true" maximal value which could be attained at that particular time. How closely the largest $FEV₁$ measurement approaches that true value cannot be determined; that is, there is an unobservable measurement error inherent in any such test. However, there also exists for any particular test an observed variability associated with each forced expiratory effort, which we have indexed by the difference between the two largest $FEV₁$ measurements at each survey for each subject $(FEV_1$ difference). Underlying the use of a repeatability criterion is the assumption that the observed variation between blows reflects the unobservable error of the measurement; thus, a maximum allowable FEV₁ difference (e.g., 200 ml) can be used to identify, and therefore exclude, measurements which are most likely to underestimate substantially the true $FEV₁$. We have found, however, that a nonrepeatable test may reflect not only measurement error but also pathology in the tested subject. The findings reported here suggest that the exclusion of study subjects without repeatable pulmonary function tests may result in selection bias.

We have reviewed some of the epidemiologic studies of pulmonary function reported since the recommendations of the Snowbird Workshop. In crosssectional studies, the proportion of subjects excluded from analysis because of the unsatisfactory repeatability of their tests was usually less than 10 per cent, and often less than 5 per cent (3, 5, 7).

Evidence from our longitudinal study suggests that some of those subjects with a test failure would have been likely to fail again if retested and, moreover, that these individuals would on average lose $FEV₁$ at a faster rate than those who had repeatable tests. The effects of excluding a small percentage of subjects in a crosssectional study of an environmental pulmonary hazard would depend on the prevalence of pulmonary disease in the study population. If the prevalence of disease were high in relationship to a particular exposure, and if the study population were sufficiently large, the exclusion of a small percentage of the group because of test failures would be unlikely to influence greatly the overall exposure-response relationship, even if all those excluded had lung disease. On the other hand, if the prevalence of disease were low (less than 5 per cent), the exclusion of even a few subjects with disease related to the environmental exposure might result in no significant association between exposure and disease being observed. This bias to the null may be particularly important for occupational settings in which pulmonary disease may be an idiosyncratic response, such as a specific allergy causing occupational asthma. In such a setting, self-selection of affected workers away from hazardous exposures frequently occurs. If we then further exclude affected individuals because of poor repeatability, the likelihood of finding a pulmonary effect may be reduced substantially.

In recently published longitudinal studies $(4, 6, 8-10)$ (table 4), the proportions of subjects excluded for test failures were higher than in cross-sectional studies because inclusion was often made conditional upon having repeatable pulmonary function tests in each of several surveys. An extreme example is a recent study of air pollution (9), in which the authors excluded 41 per cent of their study population by requiring a "completely ac-

Study (reference no.)	No. of surveys	Repeatability rule: two largest efforts must be	% excluded by the rule	
2 Such that observer felt maximal Bosse et al. (4) effort had been made in both surveys		$25*$		
Diem et al. (6)	9	Within 3% in at least three of the nine surveys	20	
Jones et al. (8)	2	Within 3% in both surveys	10†	
Van der Lende et al. (9)	4	Within 100 ml for FEV ₁ and within 150 ml for forced vital capacity in all four surveys	41	
Wegman et al. (10)	з	Within 5% or 100 ml in all sur- veys	8	

TABLE 4 *Selected studies of pulmonary function reported after standardization report of the Snowbird Workshop*

* No differences were found between those excluded and the rest of the cohort with respect to age or percentage of current smokers.

t This 10% includes subjects excluded because of a poor start or failure to maintain effort, and indicates poor repeatability.

ceptable lung function reading all four times (e.g., maximum effort according to technician, shape curve satisfactory, difference among three VC readings no more than 150 ml, among three FEV_1 readings no more than 100 ml)." The impact of repeatability restrictions is therefore potentially greater in longitudinal studies.

The exclusion of such a large number of subjects is avoidable. In our study, for example, by requiring that repeatable $FEV₁s$ be present in just two of the six surveys, only 3 per cent of the survivor cohort were excluded. Although one may choose to exclude nonrepeatable *measurements* from the estimation of the rates of decline in lung function, *subjects* should not be excluded from the study as long as they have two or more repeatable tests that permit their rate of decline to be estimated. Investigators should also consider the order of occurrence of test failures and successes within individuals and examine whether estimates of lung function decline could be systematically affected by the sequence of failures and successes.

Editor's note. The findings in the preceding paper by Eisen et al. have somewhat broader application than may be apparent at first glance. In addition to other measures of ventilatory function, they probably apply to periodic tests of any characteristic which require maximal physical or mental effort, such as grip strength; Master's two-step test and its modern successor, the treadmill; or tests of memory and calculating ability.

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